Análisis Secuencial Bayesiano de Diseños Experimentales Discriminantes Definitivos

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Credits

■ Joint work with Román de la Vara (CIMAT).



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- If m number of factors, require 2m + 1 experimental runs
- There are potentially $m + \frac{m(m-1)}{2} + m + 2$ parameters!
- *m* = 6 implies 13 runs and 29 parameters. Effect sparsity and effect heredity assumptions are required for analysis

A (1) > (1) > (1)

Experimental Plan for m=6

Run	$x_1(A)$	$x_2(B)$	$x_3(C)$	$x_4(D)$	$x_5(E)$	$x_6(F)$
1	0	1	$^{-1}$	$^{-1}$	-1	-1
2	0	-1	1	1	1	1
3	1	0	-1	1	1	-1
4	-1	0	1	-1	-1	1
5	-1	-1	0	1	-1	-1
6	1	1	0	-1	1	1
7	-1	1	1	0	1	-1
8	1	-1	-1	0	-1	1
9	1	-1	1	-1	0	-1
10	-1	1	-1	1	0	1
11	1	1	1	1	-1	0
12	-1	-1	-1	-1	1	0
13	0	0	0	0	0	0

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Model

$$\begin{aligned} y_i &= \beta_0 + \sum_{j=1}^m \beta_j x_{ij} + \sum_{j=1}^{m-1} \sum_{k=j+1}^m \beta_{jk} x_{ij} x_{ik} + \sum_{j=1}^m \beta_{jj} x_{ij}^2 + \varepsilon_i \\ i &= 1, ..., 2m + 1 \\ \varepsilon_i \text{ iid } N(0, \sigma^2) \\ \frac{(m+2)(m+1)}{2} + 1 \text{ unknown parameters: } \beta, \sigma^2 \end{aligned}$$

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JN11 suggested for Analysis

Forward step regression programmed in JMP

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JN11 suggested for Analysis

- Forward step regression programmed in JMP
- Model search based on a bias corrected version of Akaike's Information Criteria

Forward Stepwise Regression in JMP

■ Follows strategy by Hamada and Wu (1992)



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Forward Stepwise Regression in JMP

- Follows strategy by Hamada and Wu (1992)
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Forward Stepwise Regression in JMP

- Follows strategy by Hamada and Wu (1992)
- Iterative procedure. JN11 used the "Combine" option with a p-value to enter of 0.1
- Our strategy may be viewed as a adaptation of Hamada and Wu's approach to Definitive Screening Experiments and using Bayesian tools

 This procedure is based on a bias correction for Akaike's Information Criteria given in Hurvich and Tsai (1989)



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- Let n = 2m + 1 and k(< n) be the number of parameters in β in a given model M

$$AIC_{c} = n\ln(\widehat{\sigma}^{2}) + n\frac{1+m/n}{1-(m+2)/n}$$

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- JN mention that with n = 13, 10 is the maximum number of predictors that can be included with AICc
- JN11 fitted all possible models having 10 or fewer predictors, 16, 628, 808 models

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- The joint posterior distribution of β is multivariate t with $\nu = n k$ degrees of freedom
- **Then the posterior distribution of each** β_i is

$$eta_j | \mathbf{y}$$
, $m{M} \sim t(\widehat{eta}_j, s^2 c_{jj},
u)$

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Bayesian Evaluation of Significance

Posterior probability intervals $P(a < \beta_i < b | \mathbf{y}, M) = 1 - \alpha$



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Bayesian Evaluation of Significance

- Posterior probability intervals $P(a < \beta_i < b | \mathbf{y}, M) = 1 \alpha$
- Odds that the parameter is positive or negative using $P(\beta_i > 0 | \mathbf{y}, M)$ or $P(\beta_i < 0 | \mathbf{y}, M)$

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- Odds that the parameter is positive or negative using $P(\beta_i > 0 | \mathbf{y}, M)$ or $P(\beta_i < 0 | \mathbf{y}, M)$
- Note: Care has to be taken so that k < n, up to k = n 1

1 A an event: odds in favor
$$\frac{P(A)}{1-P(A)}$$
.; odds against $\frac{1-P(A)}{P(A)}$



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- **1** A an event: odds in favor $\frac{P(A)}{1-P(A)}$; odds against $\frac{1-P(A)}{P(A)}$
- Interpretation: odds in favor greater than one, then it is the number of times that A is more likely to happen than its complement.

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- If you have odds in favor 10-1 or 20-1 of winning in a game, would you bet on it?

Image: A math a math

Step 1. Fit a model with all main effects. Choose those that show some potential of significance. Be generous.



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- Step 2. Using the effects chosen in Step 1, fit a model with main, quadratic and interaction effects. Choose those that show potential of significance. Be strict.
- Step 3. Fit a model with all main effects again, and quadratic and interaction effects that showed significance in Step 2. Choose those that show potential of significance. Be strict.

Step 4. Using the effects chosen in Step 3, fit a model with main, quadratic and interaction effects that have shown significance. If any new main effect should be included add it with any quadratic or interaction effects that have not been tested before. Choose those that show potential of significance. Be strict.

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- Step 4. Using the effects chosen in Step 3, fit a model with main, quadratic and interaction effects that have shown significance. If any new main effect should be included add it with any quadratic or interaction effects that have not been tested before. Choose those that show potential of significance. Be strict.
- **Final Step**. Fit a model with effects that have shown evidence of being significance.

Simulated Example

$$m = 6$$

$$y_j = 20 + 4A + 3B - 2C - D + 5BC + 6AA + \varepsilon_i$$

 $arepsilon_i \sim N(0, 1)$

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Stepwise Procedure in JMP

Stepwise Results

Stepwise Fit for Y									
d CL	irrent	Estimates							
Los	kEnte	red Parameter	Estimate	nDF	\$\$	"F Ratio"	"Prob>F"		
		Intercept	20.0576471	1	0	0.000	1		
	1	X1	3.408	2	192,6438	407.638	2.9e-6		
	1	3/2	2.748	2	200.7311	424.752	2.62e-6		
		X3	-1.309	2	142.3509	301.218	8.15e-6		
	1	3/4	-0.851	2	9.407917	19.907	0.00416		
		X5	0	1	0.26896	1.179	0.33862		
		208	0	1	0.43264	2.311	0.20309		
		X1*X2	0	1	0.152025	0.591	0.48501		
		X1*X3	0	1	0.131951	0.503	0.51737		
		X1*X4	0	1	0.158025	0.809	0.47889		
		X1*X5	0	2	0.435212	0.875	0.50199		
		×1*×8	0	2	0.779243	2.906	0.19864		
	1	X2*X3	5.595	1	125.2161	529.920	2.88e-6		
		X2*X4	0	1	0.461888	2.568	0.19434		
		X2*X5	0	2	0.70216	2.197	0.25839		
		3(2*)(8	0	2	0.43281	0.967	0.50442		
		X3*X4	0	1	0.01918	0.068	0.80991		
		X3*X5	0	2	0.417108	0.819	0.52037		
		X3*X8	0	2	0.473473	1.003	0.46388		
		X4*X5	0	2	0.279325	0.464	0.86724		
		X4*X8	0	2	0.596892	1.542	0.34825		
		X5*X8	0	3	0.857625	1.766	0.38153		
	2	201920	-7.2714708	1	78.49896	323.747	0.00001		
		X2*X2	0	1	0.01918	0.068	0.80991		
		X3*X3	0	1	0.152025	0.591	0.48501		
	1	364*364	1.22352941	1	2.165907	9.166	0.02916		
		X5*X5	0	2	0.520167	1.180	0.41876		
_	_	20000-00			0.000047	0.004	0.07004		

JMP finds as active: A, B, C, D, BC, AA, DD. Type I error to include DD. Notice: AA reported in JMP is -7.27 (vs reported true=6)

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Minimization of Hurvich and Tsai's AICc

$$y_j = 20 + 4A + 3B - 2C - D + 5BC + 6AA + \varepsilon_i$$

Minimizing AIC_c gives model: A, B, C, BC, AA with $AIC_c = 70.63$. Type II error not to include D.



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- AIC_c = 83.72 for model found by JMP:
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- AIC_c = 83.72 for model found by JMP:
 A, B, C, D, BC, AA, DD. Type I error to include DD.
- 3 $AIC_c = 71.25$ for true model A, B, C, D, BC, AA

Step 1. Posterior Distributions



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Step 1. Posterior Probability Intervals and Odds

Effects	Estim	Q025	Q975	P+	P-	Odds +	Odds -
Α	3.408	-0.839	7.655	0.951	0.049	19.6	0.1
В	2.748	-1.499	6.995	0.918	0.082	11.2	0.1
С	-1.309	-5.556	2.938	0.24	0.76	0.3	3.2
D	-0.851	-5.098	3.396	0.321	0.679	0.5	2.1
E	-0.164	-4.411	4.083	0.464	0.536	0.9	1.2
F	-0.208	-4.455	4.039	0.454	0.546	0.8	1.2

Potentially Significant A, B, and C

Step 2. Posterior Distributions



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Step 2. Posterior Probability Intervals and Odds

Effects	Estim	Q025	Q975	P+	P-	Odds +	Odds -
Α	3.408	1.77	5.046	0.996	0.004	284.3	0.0
В	2.748	1.11	4.386	0.994	0.006	154.8	0.0
С	-1.309	-2.947	0.329	0.042	0.958	0.0	22.7
AA	-6.579	-10.557	-2.602	0.007	0.993	0.0	148.7
BB	0.323	-3.655	4.301	0.594	0.406	1.5	0.7
CC	0.4	-3.577	4.378	0.615	0.385	1.6	0.6
AB	0.722	-1.652	3.095	0.798	0.202	3.9	0.3
AC	-0.249	-2.622	2.124	0.38	0.62	0.6	1.6
BC	4.874	2.5	7.247	0.996	0.004	273.7	0.0

Potentially Significant A, B, C, AA, BC

Step 3. Posterior Distributions



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Step 3. Posterior Probability Intervals and Odds

Effects	Estim	Q025	Q975	P+	P-	Odds +	Odds -
Α	3.408	2.694	4.122	1	0	10670.0	0.0
В	2.748	2.034	3.462	1	0	4600.6	0.0
С	-1.309	-2.023	-0.595	0.004	0.996	0.0	283.4
D	-0.851	-1.565	-0.137	0.015	0.985	0.0	66.4
E	-0.164	-0.878	0.55	0.279	0.721	0.4	2.6
F	-0.208	-0.922	0.506	0.232	0.768	0.3	3.3
AA	-6.725	-8.404	-5.045	0	1	0.0	5368.4
BC	5.152	4.25	6.054	1	0	21648.4	0.0

Potentially Significant A, B, C, D, AA, BC

Step 4. Posterior Distributions



Step 4. Posterior Probability Intervals and Odds

Effects	Estim	Q025	Q975	P+	P-	Odds +	Odds -
Α	3.408	2.602	4.214	0.998	0.002	664.2	0.0
В	2.748	1.942	3.554	0.998	0.002	432.5	0.0
С	-1.309	-2.115	-0.503	0.01	0.99	0.0	99.7
D	-0.851	-1.657	-0.045	0.023	0.977	0.0	43.3
AA	-7.12	-13.982	-0.258	0.023	0.977	0.0	41.8
DD	1.264	-3.578	6.106	0.811	0.189	4.3	0.2
AD	-0.072	-2.029	1.885	0.444	0.556	0.8	1.3
BC	5.665	1.439	9.891	0.986	0.014	68.5	0.0
BD	-0.249	-1.417	0.919	0.228	0.772	0.3	3.4
CD	0.002	-2.402	2.406	0.501	0.499	1.0	1.0

Potentially Significant A, B, C, D, AA, BC

Final Step. Posterior Distributions



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Final Step. Posterior Probability Intervals and Odds

$y_j = 20 + 4A + 3B - 2C - D + 5BC + 6AA + \varepsilon_i$

Effects	Estim	Q005	Q995	P+	P-	Odds +	Odds -
Α	3.408	2.53	4.28	1	0	288050.3	0.0
В	2.748	1.87	3.62	1	0	82345.0	0.0
С	-1.309	-2.19	-0.43	0.001	0.999	0.0	1371.8
D	-0.851	-1.73	0.03	0.006	0.994	0.0	175.6
AA	-6.725	-8.78	-4.67	0	1	0.0	103581.7
BC	5.152	4.05	6.26	1	0	827895.2	0.0

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- Requires fitting a very limited set of models
- If possible, it should be used in conjunction with the other methods
- It could be adapted to other screening designs
- It is programmed in R

To obtain posterior probability intervals for a mean response: $P(a < \mathbf{x}_0^T \boldsymbol{\beta} < b | \mathbf{y}, M) = 1 - \alpha$

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- To get posterior prediction intervals $P(a < y'_0 < b | \mathbf{y}, M, \mathbf{x}_0) = 1 \alpha$
- Compare the procedures by simulation
- It may be applied to Definitive Screening Experiments when the response is non-normal (GLM). MCMC

Thank you very much!

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References

- Hamada, M. S. and Wu, C. F. J. (1992). "Analyzing Designed Experiments with Complex Aliasing". Journal of Quality Technology 24, pp. 30–37.
- Hurvich, C. M. and Tsai, C.-L. (1989). "Regression and Time Series Model Selection in Small Samples". *Biometrika* 76, pp. 297–330.
- Jones, B. and Nachtsheim, C. J. (2011). "A Class of Three-Level Designs for Definitive Screening in the Presence of Second-Order Effects" *Journal of Quality Technology* 43, pp. 1-14.